

Learning traffic correlations in multi-class queueing systems by sampling workloads

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Joint work with *Michel Mandjes* (U. of Amsterdam)

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Motivation: Backbone of the Internet

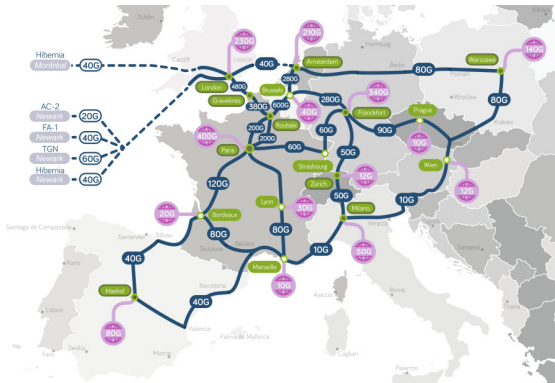


Figure: OVH Europe network

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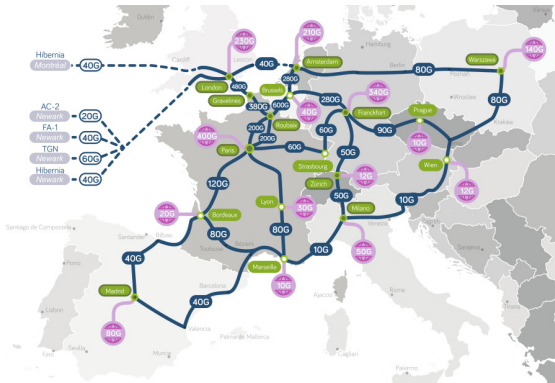


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Features:

- **Traffic:** Highly aggregated

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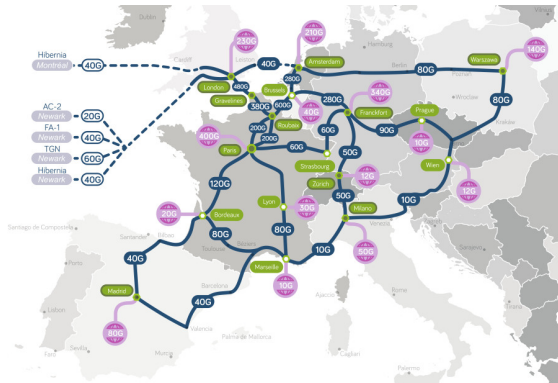
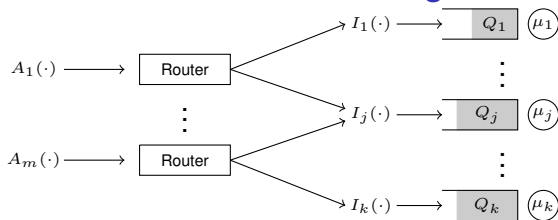


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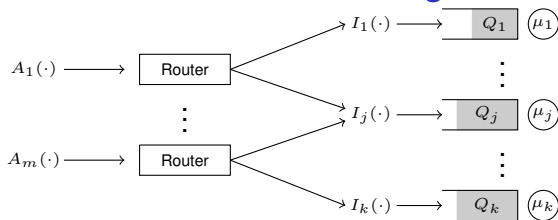
Features:

- ▶ **Traffic:** Highly aggregated
- ▶ **Routing:** Mostly static

Stylized model: Static load balancing



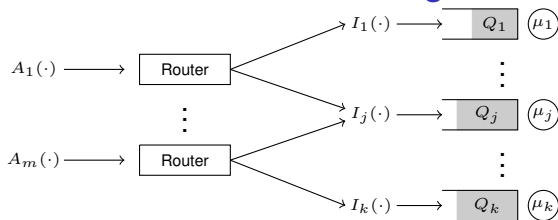
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Assumptions:

- **Arrivals:** $A(\cdot)$ is Gaussian with known rate $\lambda \in \mathbb{R}_+^m$ and unknown covariance matrix $\Sigma : \mathbb{R}^2 \rightarrow \mathbb{R}^{m \times m}$

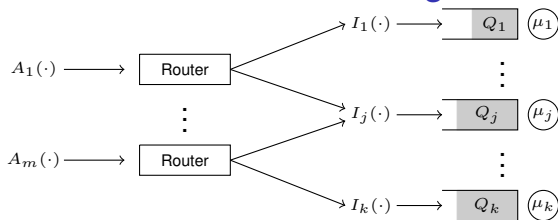
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Objective: Learn

$$R^* \in \arg \min \left\{ \max_{i \in \{1, \dots, k\}} \left\{ \mathbb{P}(Q_i(R) > b_i) \right\} \right\}$$

Additional assumptions on $A(\cdot)$

Assumption: Many-sources regime

$$A(\cdot) = A^{(n)} = \frac{1}{n} \sum_{i=1}^n X^{(i)}(\cdot),$$

where $\{X^{(i)}(\cdot)\}_{i \geq 1}$ are i.i.d.

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For simplicity: Consider multivariate fractional Brownian motions (mfBm)

$$\text{Cov} \left(A_i^{(n)}(t), A_j^{(n)}(s) \right) = \frac{\sigma_i \sigma_j \rho_{i,j}}{2n} \left(|t|^{2H} + |s|^{2H} - |s - t|^{2H} \right)$$

- ▶ Hurst parameter: $H \in (0, 1)$
- ▶ Variance: $\sigma_i^2 > 0$
- ▶ Correlation: $\rho_{i,j} \in [-1, 1]$

How do we find the optimal routing matrix?

Approach 1: directly learn model parameters

Algorithm:

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Disadvantage:

- ▶ It is hard to estimate covariance matrix

Approach 2: directly learn optimal routing

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Advantage:

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Disadvantage:

- ▶ Possibly slow to converge

Optimization with indirect learning

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Advantages:

- ▶ Queue lengths are easier to estimate than covariances
- ▶ Fast convergence

First inversion procedure

From marginal queue lengths to variances

[Mandjes & van de Meent (2009) Resource dimensioning through buffer sampling]

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Work flow:

$$A^{(n)}(\cdot) \xrightarrow{\text{Routing}} I_i^{(n)}(\cdot) \xrightarrow{\text{Queueing}} Q_i^{(n)}$$

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$$Q_i^{(n)} \xrightarrow{\text{"Inversion"}} \text{Var} \left(I_i^{(n)}(\cdot) \right)$$

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based on the large-deviations principle

$$\lim_{n \rightarrow \infty} -\frac{1}{n} \log \left(\mathbb{P} \left(Q_i^{(n)} > b \right) \right) = \inf_{t < 0} \left\{ \frac{[b - (\mu_i - \bar{\lambda}_i)t]^2}{2n \text{Var} \left(I_i^{(n)}(t) \right)} \right\}$$

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$$n \text{Var} \left(I_i^{(n)}(t) \right) \leq \frac{[b - (\mu_i - \bar{\lambda}_i)t]^2}{-2 \lim_{n \rightarrow \infty} \frac{1}{n} \log \left(\mathbb{P} \left(Q_i^{(n)} > b \right) \right)}$$

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$$\text{Var} \left(I_i^{(n)}(t) \right) \lesssim \inf_{b \in [\epsilon, 1/\epsilon]} \left\{ \frac{[b - (\mu_i - \bar{\lambda}_i)t]^2}{-2 \log \left(\hat{\mathbb{P}}_N \left(Q_i^{(n)} > b \right) \right)} \right\}$$

From marginal queue lengths to variances

Variance estimator:

$$\hat{V}_{i,\epsilon}^{(n,N)}(t) \triangleq \inf_{b \in [\epsilon, 1/\epsilon]} \left\{ \frac{[b - (\mu_i - \bar{\lambda}_i)t]^2}{-2 \log \left(\hat{\mathbb{P}}_N \left(Q_i^{(n)} > b \right) \right)} \right\} \stackrel{?}{\approx} \text{Var} \left(I_i^{(n)}(t) \right)$$

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Theorem

Fix $t < 0$. We have

$$\lim_{n \rightarrow \infty} \lim_{N \rightarrow \infty} n \left| \hat{V}_{i,\epsilon}^{(n,N)}(t) - \text{Var} \left(I_i^{(n)}(t) \right) \right| = 0, \quad a.s.,$$

for all ϵ small enough.

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- ▶ Need to find the $m(m + 1)/2$ distinct $Cov \left(A_i^{(n)}(\cdot), A_j^{(n)}(\cdot) \right)$

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- ▶ For each of the k queues we obtain a linear equation

$$\hat{V}_{i,\epsilon}^{(n,N)}(\cdot) \approx Var \left(I_j^{(n)}(\cdot) \right) = \sum_{j=1}^m \sum_{q=1}^m R_{j,i} R_{q,i} Cov \left(A_j^{(n)}(\cdot), A_q^{(n)}(\cdot) \right)$$

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Use joint queue lengths to get more equations directly?

Second inversion procedure

From pair-wise joint queue lengths to covariances

Work flow:

$$A^{(n)}(\cdot) \xrightarrow{\text{Routing}} \left(I_i^{(n)}(\cdot), I_j^{(n)}(\cdot) \right) \xrightarrow{\text{Queueing}} \left(Q_i^{(n)}, Q_j^{(n)} \right)$$

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Inversion:

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Feasibility of the inversion

Covariance estimator:

$$\hat{C}_{i,j,\epsilon}^{(n,N)}(t,s) \triangleq \inf_{c_i, c_j \in [\epsilon, 1/\epsilon]} \left\{ \dots \right\} \stackrel{?}{\approx} \text{Cov} \left(I_i^{(n)}(t), I_j^{(n)}(s) \right)$$

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Theorem

If $I^{(n)}(\cdot)$ is short-range dependent ($H \leq 1/2$), and $I_i^{(n)}(\cdot)$ and $I_j^{(n)}(\cdot)$ are non-negatively correlated ($\rho_{i,j} \geq 0$), then

$$\lim_{n \rightarrow \infty} \lim_{N \rightarrow \infty} n \left| \hat{C}_{i,j,\epsilon}^{(n,N)}(t,t) - \text{Cov} \left(I_i^{(n)}(t), I_j^{(n)}(t) \right) \right| = 0, \quad a.s.,$$

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- ▶ For each of the $k(k - 1)/2$ pairs (i, j) , with $i < j$, we get

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- ▶ If $m \leq k$, these are enough for a single R

Fundamental limitation

Theorem

If $I^{(n)}(\cdot)$ is long-range dependent ($H > 1/2$), and $I_i^{(n)}(\cdot)$ and $I_j^{(n)}(\cdot)$ are negatively correlated ($\rho_{i,j} < 0$), then $Cov\left(I_i^{(n)}(t), I_j^{(n)}(t)\right)$ **cannot be recovered** from the large deviations behavior of $c_i Q_i^{(n)} + c_j Q_j^{(n)}$.

There is an inherent loss of information!

Why we have this
limitation?

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Three explanations:

- ▶ The analytical one
- ▶ Sample path large deviations interpretation
- ▶ Intuitive with queues as reflected processes

Analytical explanation

Recall:

$$-\log \left(\mathbb{P} \left(c_i Q_i^{(n)} + c_j Q_j^{(n)} > 1 \right) \right) \\ \approx \inf_{t,s < 0} \left\{ \frac{\left[1 - c_i (\mu_i - \bar{\lambda}_i) t - c_j (\mu_j - \bar{\lambda}_j) s \right]^2}{2 \text{Var} \left(c_i I_i^{(n)}(t) + c_j I_j^{(n)}(s) \right)} \right\}$$

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We show: If (t^*, s^*) is a minimizer \Rightarrow either $t^* = 0$ or $s^* = 0$, and

$$-\log \left(\mathbb{P} \left(c_i Q_i^{(n)} + c_j Q_j^{(n)} > 1 \right) \right) \\ \approx \min \left\{ \inf_{t < 0} \left\{ \frac{\left[1 - c_i (\mu_i - \bar{\lambda}_i) t \right]^2}{2 c_i^2 \text{Var} \left(I_i^{(n)}(t) \right)} \right\}, \inf_{s < 0} \left\{ \frac{\left[1 - c_j (\mu_j - \bar{\lambda}_j) s \right]^2}{2 c_j^2 \text{Var} \left(I_j^{(n)}(s) \right)} \right\} \right\}$$

Sample path large deviations interpretation

Let: $q_i(\cdot)$ and $q_j(\cdot)$ be most likely paths for the event

$$\left\{ c_i Q_i^{(n)}(0) + c_j Q_j^{(n)}(0) = 1 \right\}$$

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Note: The minimizers (t^*, s^*) are such that

$$t^* = \sup \{t < 0 : q_i(t) = 0\} \quad \text{and} \quad s^* = \sup \{s < 0 : q_j(s) = 0\}$$

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Here: Since either $t^* = 0$ or $s^* = 0$, most likely paths are in

$$\left\{ c_i Q_i^{(n)}(0) = 1 \text{ and } c_j Q_j^{(n)}(0) = 0 \right\} \cup \left\{ c_i Q_i^{(n)}(0) = 0 \text{ and } c_j Q_j^{(n)}(0) = 1 \right\}$$

Intuition with queues as reflected processes

Note: $Q_i^{(n)}(\cdot)$ and $Q_j^{(n)}(\cdot)$ are the reflection at 0 of

$$W_i^{(n)}(t) \triangleq I_i^{(n)}(t) - \mu_i t \quad \text{and} \quad W_j^{(n)}(t) \triangleq I_j^{(n)}(t) - \mu_j t$$

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In this case: Because of negative correlation:

- ▶ When $I_i^{(n)}(\cdot)$ grows faster, $I_j^{(n)}(\cdot)$ grows slower
- ▶ $Q_i^{(n)}(\cdot)$ increases and $Q_j^{(n)}(\cdot)$ decreases (until it's reflected)

Intuition with queues as reflected processes

Note: $Q_i^{(n)}(\cdot)$ and $Q_j^{(n)}(\cdot)$ are the reflection at 0 of

$$W_i^{(n)}(t) \triangleq I_i^{(n)}(t) - \mu_i t \quad \text{and} \quad W_j^{(n)}(t) \triangleq I_j^{(n)}(t) - \mu_j t$$

In this case: Because of negative correlation:

- ▶ When $I_i^{(n)}(\cdot)$ grows faster, $I_j^{(n)}(\cdot)$ grows slower
- ▶ $Q_i^{(n)}(\cdot)$ increases and $Q_j^{(n)}(\cdot)$ decreases (until it's reflected)
- ▶ Magnitude of correlation is lost in the reflection at 0

Back to finding the
optimal routing matrix

Estimating the overflow probabilities

- ▶ We have the estimates

$$\hat{\mathcal{E}}_{\epsilon, j, \ell}^{(n, N)}(t, t) \approx \text{Cov} \left(A_j^{(n)}(t), A_\ell^{(n)}(t) \right),$$

Estimating the overflow probabilities

- ▶ We have the estimates

$$\hat{\mathcal{C}}_{\epsilon,j,\ell}^{(n,N)}(t,t) \approx \text{Cov} \left(A_j^{(n)}(t), A_\ell^{(n)}(t) \right),$$

- ▶ For any matrix R , we approximate $\mathbb{P} \left(Q_i^{(n)}(R) > b_i \right)$ as

$$\hat{P}_{\epsilon,\delta,i}^{(n,N)}(R) \triangleq \exp \left(- \inf_{t \in [-1/\delta, -\delta]} \left\{ \frac{\left[b_i - \left(\mu_i - \sum_{j=1}^m R_{j,i} \lambda_j \right) t \right]^2}{\sum_{j=1}^m \sum_{\ell=1}^m R_{j,i} R_{\ell,i} \hat{\mathcal{C}}_{\epsilon,j,\ell}^{(n,N)}(t,t)} \right\} \right)$$

Estimating optimal routing matrix

Estimate:

$$\hat{R}_{\epsilon, \delta}^{(n, N)} \in \arg \min_{R \in \mathcal{R}} \left\{ \max_{i \in \{1, \dots, k\}} \left\{ \hat{P}_{\epsilon, \delta, i}^{(n, N)}(R) \right\} \right\}$$

Estimating optimal routing matrix

Estimate:

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Theorem

For any finite set of routing matrices \mathcal{R} , we have

$$\lim_{n \rightarrow \infty} \lim_{N \rightarrow \infty} \frac{1}{n} \log \left(\frac{\max_{i \in \{1, \dots, k\}} \left\{ \mathbb{P} \left(Q_i^{(n)} \left(\hat{R}_{\epsilon,\delta}^{(n,N)} \right) > b_i \right) \right\}}{\min_{R \in \mathcal{R}} \left\{ \max_{i \in \{1, \dots, k\}} \left\{ \mathbb{P} \left(Q_i^{(n)}(R) > b_i \right) \right\} \right\}} \right) = 0, \quad a.s.,$$

for all ϵ and δ small enough.

Simulations

Simulations' setup

- ▶ $n = 1000$

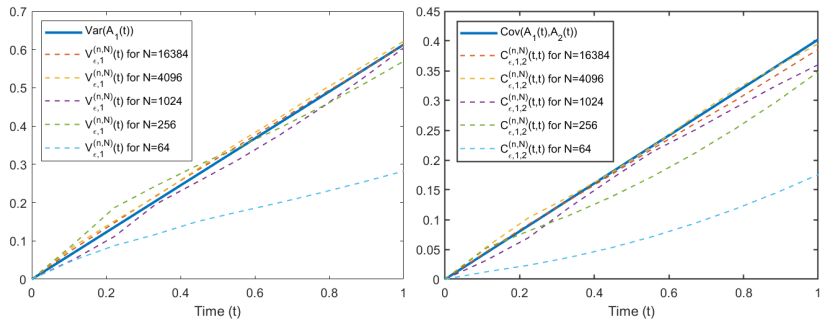
Simulations' setup

- ▶ $n = 1000$
- ▶ $m = k = 2$ and $R = Id$

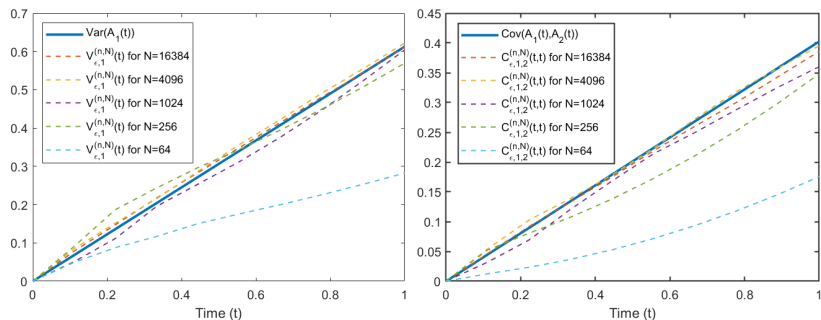
Simulations' setup

- ▶ $n = 1000$
- ▶ $m = k = 2$ and $R = Id$
- ▶ $\lambda_1 = \lambda_2 = 0.8$
- ▶ $\mu_1 = \mu_2 = 1$
- ▶ $\epsilon = 0.01$

Positively correlated Brownian motions

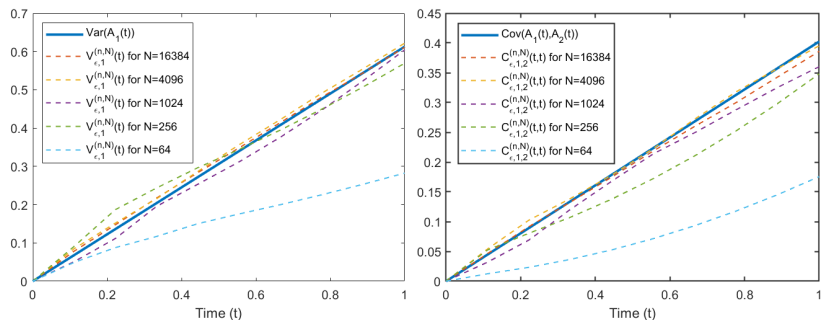


Positively correlated Brownian motions



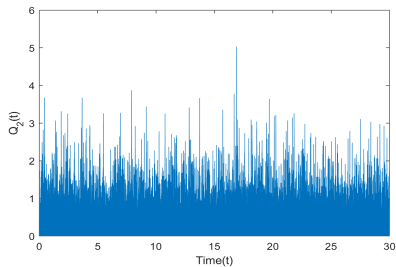
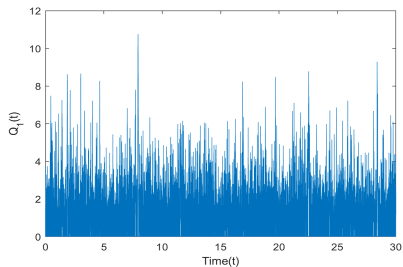
- Variance estimator works with $N \geq 256$ samples

Positively correlated Brownian motions

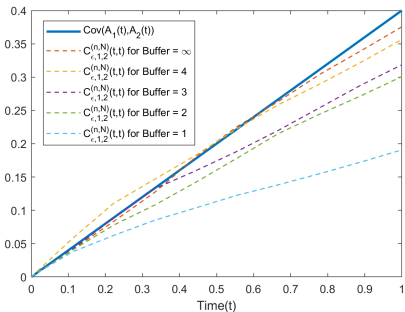
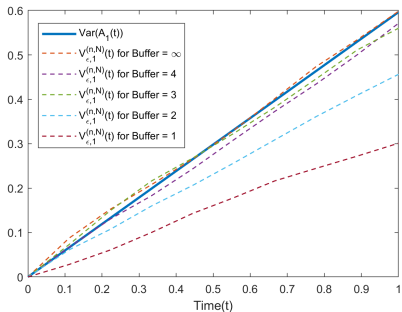
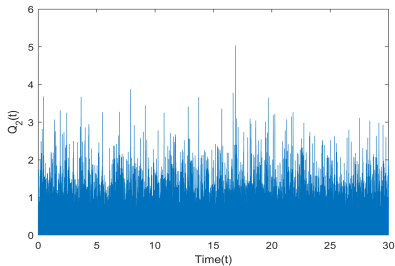
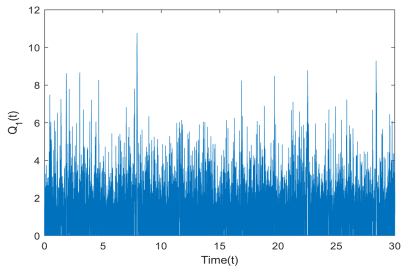


- ▶ Variance estimator works with $N \geq 256$ samples
- ▶ Covariance estimator works with $N \geq 1024$ samples

Positively correlated Brownian motions



Positively correlated Brownian motions



Positively correlated Brownian motions

Why does it work so well with finite buffers?

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Positively correlated Brownian motions

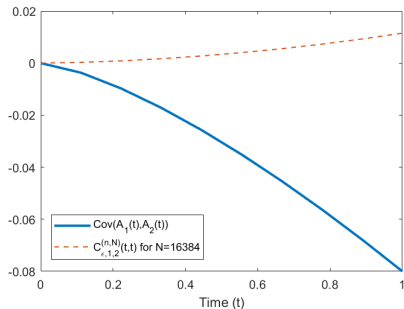
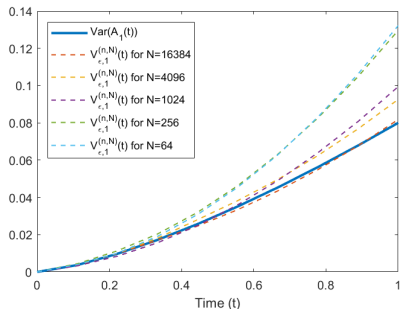
Why does it work so well with finite buffers?

- ▶ Finite buffer truncates steady-state distribution

But:

- ▶ Large-deviation approxs do not depend on buffer size
- ▶ Estimates depend on an infimum, not whole distribution

Negatively correlated fractional Brownian motions



- ▶ Variance estimator works as expected
- ▶ Covariance estimator fails as expected

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Conclusions

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⇒ Arrivals' covariances can be recovered from the queue lengths
- ▶ Long-range dependent inputs + negative correlations
⇒ Arrivals' covariances **cannot** be recovered from the queue lengths
- ▶ Can be extended to multi-path routing in acyclic networks
(needs much more involved large-deviations results)

Thank you!